Design of Two-Stage Experiments with an Application to Spillovers in Tax Compliance

Guillermo Cruces, *U. of Nottingham & CEDLAS-UNLP*Dario Tortarolo, *U. of Nottingham & IFS*Gonzalo Vazquez-Bare, *UC Santa Barbara*Julian Amendolaggine, *CEDLAS-UNLP*Juan Luis Schiavoni, *CEDLAS-UNLP*

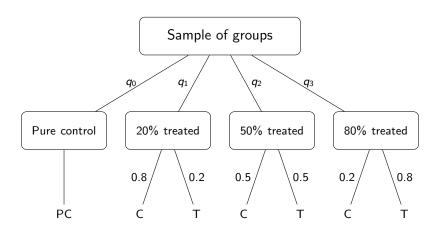
UCSB Applied Micro Economics Lunch

February 11, 2022

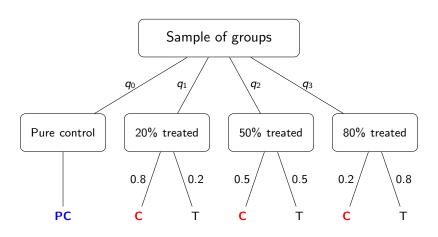
Design of Partial Population Experiments

- Goal: estimate within-group spillovers
 - ► Households in villages
 - Employees in firms
 - Students in schools
- Two-step design:
 - Groups randomly divided into treatment "intensities" (saturations)
 - Units within each group randomly assigned to treatment and control
- Compare units across groups with different treatment intensities

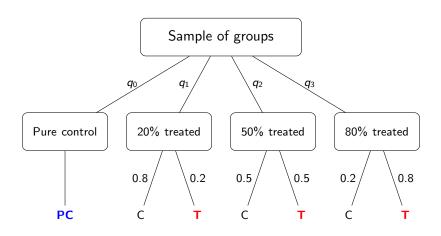
Experimental design: example



Experimental design: example



Experimental design: example



Designing PP Experiments

- Key choices:
 - ▶ Number of saturations and within-group probabilities
 - Probability of each saturation $q_0, q_1, q_2, ...$ (this talk)
 - Within-group assignment mechanism (this talk)
- Key inputs:
 - ▶ Parameters (outcome variances, intracluster correlations,...)
 - ► Variance of estimators (this talk)
 - Power function to calculate power, MDE (this talk)

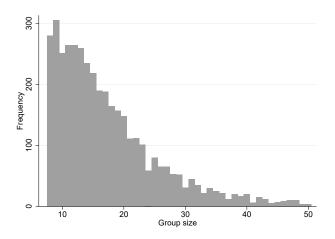
Challenges for Designing PP Experiments

- Two-stage design
- Multiple treatments
 - Compare units exposed to different saturations
- Within-group correlations (clustering)
- Heterogeneity in group sizes
 - Group sizes tend to vary widely in practice

Existing tools for designing PP Experiments

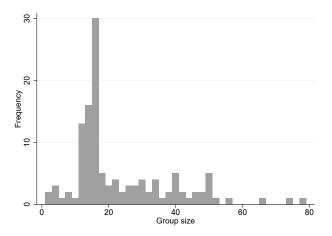
- Hirano and Hahn (2010), Baird et al (2018)
 - ► Homoskedasticity, random effects structure
 - Ignore group size heterogeneity
- Software (e.g. Stata's power command) makes restrictive assumptions about group size distribution
 - **Equally-sized groups,** N_T proportional to N_C ,...

Cruces et al (2022)



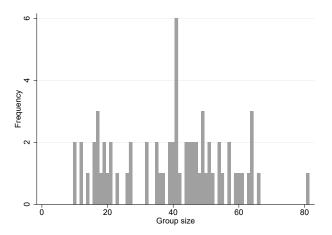
Distribution of group sizes

Haushofer and Shapiro (2016, QJE)



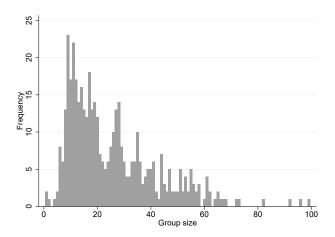
Distribution of group sizes

Giné and Mansuri (2018, AEJ Applied)



Distribution of group sizes

Imai et al (2020, JASA)

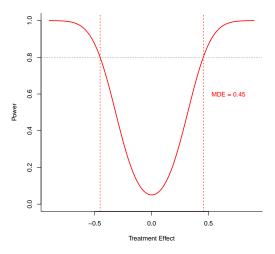


Distribution of group sizes

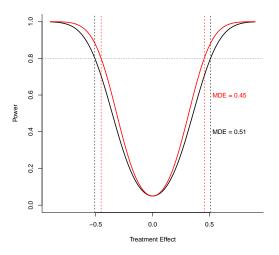
It affects the variance of estimators

$$\mathbb{V}[\hat{\beta}] \approx \sigma^2 \left[1 + \rho(ICC, \bar{n}, Var(n_g))\right]$$

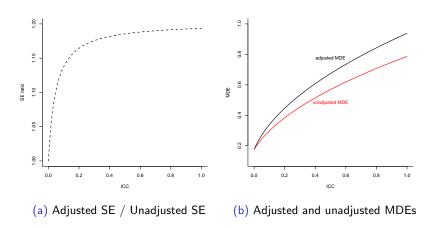
- ▶ Ignoring $Var(n_g)$ underestimates $\mathbb{V}[\hat{\beta}]$ → overestimates power
- It affects inference and power calculations
 - Normal approx may be inaccurate if groups are "too heterogeneous"
 - ► Carter et al (2017), Djogbenou et al (2019), Hansen and Lee (2019)



$$G = 95$$
, $\bar{n} = 23.3$, $sd(n_g) = 0$, $\sigma_Y^2 = 1$, $ICC = 0.2$



$$G=95,\; \bar{n}=23.3,\; sd(n_g)=15.2,\; \sigma_Y^2=1,\; ICC=0.2$$



This paper

- We derive asymptotic variance approximations allowing for:
 - Multiple treatments
 - General intracluster correlation and heteroskedasticity
 - Group size heterogeneity
 - Varying probabilities across groups
- Calculate power and MDEs
- Our formulas can be applied in a wide range of designs
 - ► Two-stage, PP, clustered, stratified experiments...
- We conduct a field experiment on tax compliance in Argentina

Setup

- Random sample of groups $g=1,\ldots,G$ with units $i=1,\ldots,n_g$
- Total sample size $n = \sum_{g} n_{g}$
- First stage: randomly divide groups into categories:

$$T_g \in \{0, 1, 2, \dots, M\}, \quad \mathbb{P}[T_g] = q_t$$

Within each group, assign binary individual-level treatment:

$$D_{ig} \in \{0,1\}, \quad \mathbb{P}_{g}[D_{ig} = d | T_{g} = t] = p_{g}(d,t)$$

Setup

Estimands:

$$\beta_{dt} = \mathbb{E}[Y_{ig}|D_{ig} = d, T_g = t] - \mathbb{E}[Y_{ig}|T_g = 0]$$

- ▶ Direct effects = β_{1t}
- ▶ Spillover effects = β_{0t}
- Second moments:

$$\begin{split} \sigma_{dt}^2 &= \mathbb{V}[Y_{ig}|D_{ig} = d, T_g = t]\\ \rho_{dt} &= cor(Y_{ig}, Y_{jg}|D_{ig} = d, D_{jg} = d, T_g = t) \end{split}$$

Setup

• Estimation strategy:

$$Y_{ig} = \alpha + \sum_{t=1}^{M} \beta_{0t} (1 - D_{ig}) \mathbb{1}(T_g = t) + \sum_{t=1}^{M} \beta_{1t} D_{ig} \mathbb{1}(T_g = t) + \varepsilon_{ig}$$

• Equivalent to:

$$\hat{\beta}_{dt} = \bar{Y}_{dt} - \bar{Y}_{00}$$

Allow for correlated errors within groups

Main Result

Asymptotic Approximation

Under regularity conditions, if

$$\max_{g \leq G} \frac{n_g^2}{n} \rightarrow 0, \quad \frac{\sum_{g=1}^G n_g^4}{n^2} \leq C < \infty,$$

then:

$$\hat{\beta}_{dt} \stackrel{a}{\sim} \mathcal{N}(\beta_{dt}, V_{dt})$$

where:

$$V_{dt} = \frac{\sigma_{dt}^{2}}{q_{t} \sum_{g} n_{g} p_{g}(d, t)} \left\{ 1 + \rho_{dt} \frac{\sum_{g} n_{g}(n_{g} - 1) \mathbb{P}_{g}[D_{ig} = d, D_{jg} = d | T_{g} = t]}{\sum_{g} n_{g} p_{g}(d, t)} \right\} + \frac{\sigma_{00}^{2}}{q_{0} n} \left\{ 1 + \rho_{00} \left(\frac{\sum_{g} n_{g}^{2}}{n} - 1 \right) \right\}$$

Main result: intuition

- Variance: $\mathbb{V}[\hat{eta}_{dt}] = \mathbb{V}[ar{Y}_{dt}] + \mathbb{V}[ar{Y}_{00}]$ allowing for:
 - ▶ Heteroskedasticity: $\sigma_{dt}^2 \neq \sigma_{d't'}^2$
 - ▶ Intracluster correlation: $\rho_{dt} \neq 0$
 - ▶ Unequal probabilities between groups: $p_g(d,t) \neq p_{g'}(d,t)$
 - Group size heterogeneity: $Var(n_g) \neq 0$

Main Result: Intuition

Condition:

$$\max_{g \le G} \frac{n_g^2}{n} \to 0$$

restricts the relative size of the largest group

- ► Ensures that no group "dominates" the sample
- Condition:

$$\frac{\sum_{g=1}^{G} n_g^4}{n^2} \le C < \infty$$

bounds the fourth moment of the distribution

Rules out fat tails (outliers)

Power and MDE calculations

• Based on the normal approximation, the power function is

$$\Gamma(\beta_{dt}) \approx 1 - \Phi\left(\frac{\beta_{dt}}{\sqrt{V_{dt}}} + z_{1-\alpha/2}\right) + \Phi\left(\frac{\beta_{dt}}{\sqrt{V_{dt}}} - z_{1-\alpha/2}\right)$$

- Depends on:
 - ▶ Treatment effect β_{dt}
 - Group sizes $\{n_g\}_{g=1}^G$ and total sample size n
 - ▶ Assig mech: $\{q_t\}_t$, $\{p_g(d,t)\}_{t,g}$, $\{\mathbb{P}_g[D_{ig}=d,D_{jg}=d|T_g=t]\}_{t,g}$
 - Outcome moments $\{\sigma_{dt}^2, \rho_{dt}\}_t$

Choice of $\{q_t\}_t$

- Optimal choice requires defining an optimality criterion
 - How to combine variances of multiple estimators
- Optimal design literature has proposed several alternatives
- We discuss two scenarios:
 - Unconstrained designs: minimize the average of all estimator variances (A-optimality)
 - Constrained designs

Choice of $\{q_t\}_t$: unconstrained optimization

A-optimal design

The solution to the optimal design problem:

$$\min_{q_0,q_1,\dots,q_M} \sum_{t=1}^M \left\{ \mathbb{V}[\hat{\beta}_{0t}] + \mathbb{V}[\hat{\beta}_{1t}] \right\}, \quad q_t > 0, \quad \sum_{t=0}^M q_t = 1$$

is:

$$q_0^* = \frac{\sqrt{2MB_0}}{\sqrt{2MB_0} + \sum\limits_{t>0} \sqrt{B_t}}, \quad q_t^* = \frac{\sqrt{B_t}}{\sqrt{2MB_0} + \sum\limits_{t>0} \sqrt{B_t}}, \quad t>0,$$

where $\{B_t\}_t$ are constants depending on $\{n_g\}_g$, $\{p_g(d,t)\}_{d,t,g}$ and $\{\mathbb{P}_g[D_{ig}=d,D_{jg}=d|T_g=t]\}_{t,g}$, $\{\sigma^2_{dt},\,\rho_{dt}\}_t$

Choice of $\{q_t\}_t$: incorporating constraints

- ullet Researchers may need to incorporate constraints in choice of q_t
 - ► Logistical, administrative, etc
- We provide an example in our field experiment
 - "Minimax-like" approach with fixed number of treated

Within-group treatment assignment

- ullet We want to assign exactly $n_g p_t$ units to treatment
- But $n_g p_t$ may not be an integer (e.g. $p_t = 0.5$, $n_g = 11$)
- ullet Let $\xi_{m{g}} \in \{0,1\}$ be a random adjustment factor and let

$$N_g^1 = \lfloor n_g p_t \rfloor + \xi_g \mathbb{1}(n_g p_t \notin \mathbb{N})$$

be the (random) number of treated in group g with $T_g = t$

• Setting $\mathbb{P}_g[\xi_g=1|T_g=t]=(n_gp_t-\lfloor n_gp_t \rfloor)\mathbb{1}(n_gp_t\notin\mathbb{N})$ gives:

$$\mathbb{E}[N_g^1|T_g=t]=n_g p_t, \quad \mathbb{P}_g[D_{ig}=1|T_g=t]=p_t$$

Direct and spillover effects in tax compliance

- We teamed up with a large municipality in Greater Buenos Aires
- Neighbors are required to pay a monthly bill on their real estate
- Information campaign with personalized letters
 - One-page letter informing of new electronic billing option
 - Instructions on how to sign up and pay online
 - Information on current billing period and past due debt
- Are there spillovers between neighbors from the same block?

Example of the intervention letter



ID: XXXXX TITULAR: DIRECCIÓN: CAP MADARIAGA Nº LOCALIDAD: 11 de Septiembre C.P.: 1657 PARTIDA: XXXXXXX/7 De esta manera, nos cuidamos entre todos al reducir la circulación u también cuidamos el medio ambiente. Es una PARTIDA: XXXXX/7 Cuota 10 vencimiento 10 de octubre 2020: 347,29 Deuda año en curso*: 1.702.58 Deuda años anteriores*: 289,54 JOÓMO PAGAR? Ingresando a tasas.tresdefebrero.gov.ar completá los datos: 1) Podés pagar ONLINE con DESCARGÁ O PAGÁ TU BOLETA ado - En el momento desde nuestra web. Obteniendo el código de pago electrónico noro popar desde la plataforma de tu banco o caiero automático. 2) Podés papar en EFECTIVO en CLICKEÁ ESTE BOTÓN apipago - DESCARGALA o levá tu NÚMERO DE PARTIDA. miboleta tresdefebrero.apv.ar

Por dudas comunicate con nosotros a **reclamos.mistasas@tresdefebrero.gov.ar**

¡Muchas gracias!

PP Experiment: Design

- We randomly divide blocks into four categories:
 - $ightharpoonup T_g = 0$: pure controls with prob q_0
 - $ightharpoonup T_g=1$: 20% treated with prob q_1
 - $ightharpoonup T_g = 2$: 50% treated with prob q_2
 - $ightharpoonup T_g = 3$: 80% treated with prob q_3
- ullet We set up a system of eqs to incorporate constraints on $\{q_t\}_t$

Constrained choice of $\{q_t\}_t$

- Choose q_1, q_2, q_3 , with $q_0 = 1 q_1 q_2 q_3$
- The total number of letters sent (*L*) should equal the expected number of treated:

$$L = n(0.2q_1 + 0.5q_2 + 0.8q_3)$$

- ullet Categories $T_g=1$ and $T_g=3$ are symmetric, so $q_1=q_3$
- This leaves two probabilities to be determined: q_2 and q_3
- Idea: balance variances across assignments

Constrained choice of $\{q_t\}_t$

- ullet The "hardest" effects (smallest cells) to estimate are eta_{03} and eta_{11}
 - ▶ Spillover effect in 80% groups and direct effect in 20% groups
- We choose q_2 and q_3 by setting:

$$\mathbb{V}[\hat{\beta}_{03}] = \mathbb{V}[\hat{\beta}_{02}]$$

based on our variance approximation

• We assume $\sigma^2=0.25$ (upper bound for binary outcomes) and $\rho\approx 0.1$ (based on baseline data)

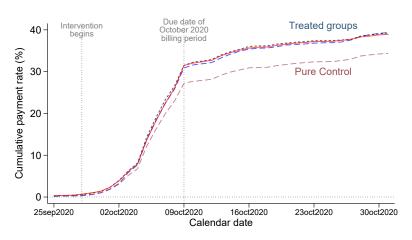
Sample sizes

	Blocks	Control Obs	Treated Obs
$T_g = 0$	1, 102	19, 105	0
$T_g = 1$	1, 100	15,049	3,864
$T_g = 2$	680	5 , 898	5 , 904
$T_{g} = 3$	1,100	3 , 707	15, 281
Total	3,982	43,759	25,049

MDEs range from 2.6 to 3.3 p.p.

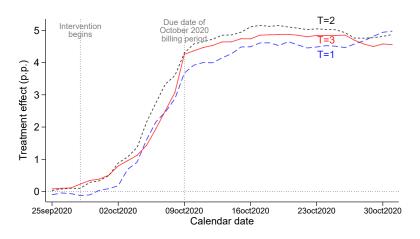
Treated groups: Payment rate (Oct'20 bill)

Figure: Payment rates in levels



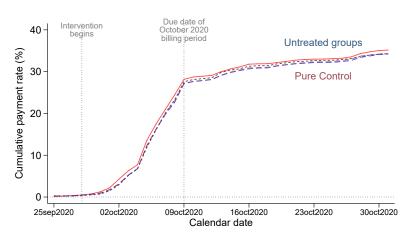
Treated groups: Payment rate (Oct'20 bill)

Figure: Difference relative to pure control group



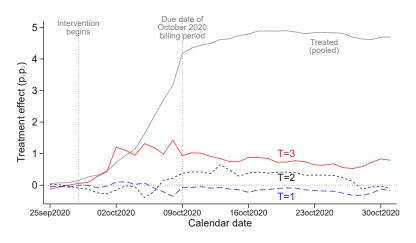
Untreated groups: Payment rate (Oct'20 bill)

Figure: Payment rates in levels

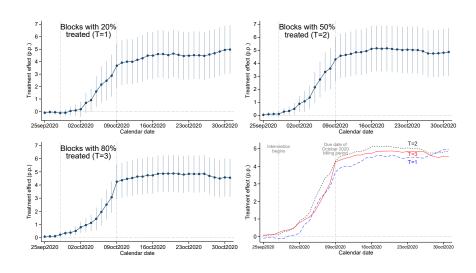


Untreated groups: Payment rate (Oct'20 bill)

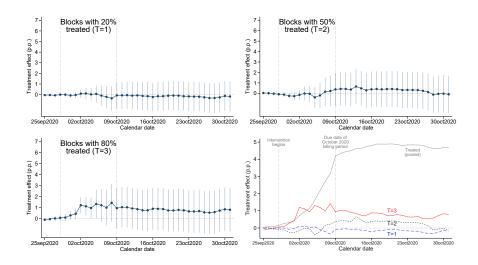
Figure: Difference relative to pure control group



Diff treated/pure controls (Oct'20 bill)



Diff. untreated/pure controls (Oct'20 bill)



Summary

- Framework to calculate power and MDE in PP experiments
 - ▶ Allow for group size heterogeneity, heteroskedasticity, ICC,...
 - Derive optimal choice of group-level probabilities
- Application to tax compliance in Argentina
 - Strong and significant direct effects of the letters
 - ▶ No clear evidence of reinforcement effects between treated
 - ▶ Some evidence of within-neighbor spillovers in highest saturation

Thank you!